

# Continuous Time Markov Chains (CTMC)

## Lecture #8

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## 1 Markov Process (Continuous Time Markov Chain)

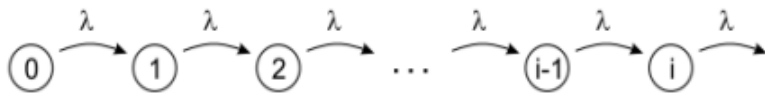
### 1.1 Transient Analysis of CTMC

Generally, transient analysis is a very challenging task. However, instantaneous state probability can be driven for a few case especially when the state probabilities at time 0,  $\pi_i(0)$ , are known. E.g. usually one knows that the system at time 0 is precisely in a given state k. The objective of the transient analysis is to find how the state probabilities evolve as a function of time  $\pi(t)$

The transient dynamics of CTMC are mainly defined by

$$\frac{d}{dt}\pi(t) = \pi(t)Q \tag{1}$$

#### 1.1.1 Pure Birth Transient Analysis



It is assumed that the system starts with zero population, that is to say  $\pi_0(0) = 1$  and  $\pi_i(0) = 0 \forall i \neq 0$ .

Using eq(1), one can write

$$\frac{d}{dt}\pi_0(t) = -\lambda\pi_0(t)$$

$$\frac{d}{dt}\pi_i(t) = -\lambda\pi_i(t) + \lambda\pi_{i-1}(t)$$

Solving these differential equations using laplace transform, we have

$$s\pi_o^*(s) - \pi_o(0) = -\lambda\pi_o^*(s) \rightarrow \pi_o^*(s) = \frac{1}{s + \lambda}$$

$$\pi_i^*(s) = \frac{\lambda}{s + \lambda} \pi_{i-1}^*(s) = \frac{\lambda^k}{(s + \lambda)^{k+1}}$$

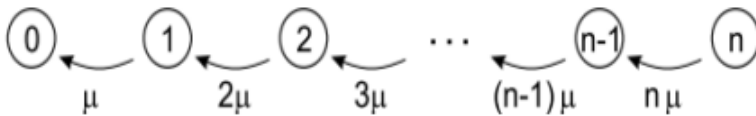
Applying laplace transform, one can express the instantensous probabilitas as

$$\pi_o(t) = e^{-\lambda t}$$

$$\pi_i(t) = \frac{(\lambda t)^i}{i!} e^{-\lambda t}$$

Note that the obtained solution indicates that the population in some period  $t$  follows a Poisson process.

### 1.1.2 Pure death Transient Analysis



It is assumed that the system starts with a population of  $n$  members, that is to say  $\pi_n(0) = 1$  and  $\pi_i(0) = 0 \forall i \neq n$

Using eq(1), one can write

$$\frac{d}{dt} \pi_n(t) = -n\mu\pi_n(t) \Rightarrow \pi_n(t) = e^{-n\mu t}$$

$$\frac{d}{dt} \pi_i(t) = (i+1)\mu\pi_{i+1}(t) - i\mu\pi_i(t) \Rightarrow \frac{d}{dt} (e^{i\mu t} \pi_i(t)) = (i+1)\mu\pi_{i+1}(t) e^{i\mu t}$$

Hence, one can express  $\pi_i(t)$  as

$$\pi_i(t) = (i+1)\mu e^{-i\mu t} \int_0^t \pi_{i+1}(\tau) e^{i\mu \tau} d\tau$$

Solving this equation for  $i = n - 1$ , we have

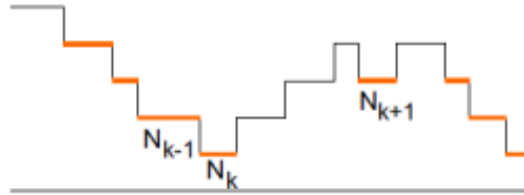
$$\begin{aligned} \pi_{n-1}(t) &= n\mu e^{-(n-1)\mu t} \int_0^t e^{-n\mu \tau} * e^{(n-1)\mu \tau} d\tau \\ &= n\mu e^{-(n-1)\mu t} \int_0^t e^{-\mu \tau} d\tau \\ &= n [(e^{-\mu t})^{n-1}] [1 - e^{-\mu t}] \end{aligned}$$

Solving recursively for  $n - 2, n - 3, \dots$  one can express the transient state probabilities as

$$\pi_i(t) = \binom{n}{i} [(e^{-\mu t})]^i [1 - e^{-\mu t}]^{n-1}$$

## 2 M/G/1 Queue Length Analysis

- The analysis is based on Embedded Markov chain or jump chain
- Let  $N_k$  be the queue length after the departure of customer  $k$
- Let  $V_k$  be the number of new customers arrived during the service time of customer  $k$ .



- Markov chain is constituted by the queue left by an departing customer.
    - Given  $N_k$ ,  $N_{k+1}$  can be expressed in terms of it and of a random variable  $V_{k+1}$
- $$N_{k+1} = \begin{cases} N_k - 1 + V_{k+1} & , N_k \geq 1 \\ V_{k+1} & , N_k = 0 \end{cases}$$
- As the service times are independent and the arrivals are Poissonian, the  $V_k$  are independent of each other.
  - $V_{k+1}$  is independent of  $N_k$  and its history
  - The stochastic characterization of  $N_{k+1}$  depends on  $N_k$  but not on the earlier history  
→ Markov Process
- Let  $a_j$  denotes the probability of  $j$  arrivals between two departures, i.e.  $P(V_k = j) = a_j$
  - Let  $p_j$  denotes the length distribution probability, then

$$p_j = p_0 a_j + \sum_{i=1}^{j+1} p_i a_{j-i+1}$$

- The MGF for the number of the customers in the system can then be derived as

$$\begin{aligned} G_N(z) &= \sum_{j=0}^{\infty} z^j p_j = \sum_{j=0}^{\infty} z^j a_j p_0 + \sum_{j=0}^{\infty} \sum_{i=1}^{j+1} z^j p_i a_{j-i+1} \\ &= p_0 G_V(z) + \sum_{i=1}^{\infty} \sum_{j=i-1}^{\infty} z^j p_i a_{j-i+1} \\ &= p_0 G_V(z) + \sum_{i=1}^{\infty} p_i \sum_{k=0}^{\infty} z^{k+i-1} a_k \\ &= p_0 G_V(z) + \sum_{i=1}^{\infty} p_i z^{i-1} \sum_{k=0}^{\infty} z^k a_k \\ &= p_0 G_V(z) + z^{-1} [G_N(z) - p_0] G_V(z) \end{aligned}$$

Hence, one can conclude that

$$G_N(z) = \frac{(z-1)p_o G_V(z)}{z - G_V(z)} \quad (2)$$

Eq (2) indicates that if one can characterize the number of arrivals between two departures  $[G_v(z)]$ , one can characterize the limiting distribution of the number of the customers in the system.

- Note that  $G_N(1) = 1 = \frac{(1-1)p_o G'_V(1) + p_o G_V(1)}{1 - G'_V(1)} \rightarrow p_o = 1 - G'(1) = 1 - \rho$
- In the following, we characterize the distribution of inter-departure arrival
  - First, let us characterize the number of arrivals from a Poisson process during a random interval  $X$

$$\begin{aligned} G_V(z) &= \sum_{v=0}^{\infty} z^v P\{V = v\} \\ &= \sum_{v=0}^{\infty} z^v \int_0^{\infty} P\{V = v | S = s\} f_S(s) ds \\ &= \sum_{v=0}^{\infty} z^v \int_0^{\infty} \frac{(\lambda s)^v}{v!} e^{-\lambda s} f_S(s) ds \\ &= \int_0^{\infty} \sum_{v=0}^{\infty} \frac{(z\lambda s)^v}{v!} e^{-\lambda s} f_S(s) ds \\ &= \int_0^{\infty} e^{z\lambda s} e^{-\lambda s} f_S(s) ds \\ &= \int_0^{\infty} e^{-(1-z)\lambda s} f_S(s) ds \\ &= G_S(\lambda(1-z)) \end{aligned} \quad (3)$$

- The same result can be attained using the law of iterated expectations

$$G_V(z) = E_V[z^k] = E_S[E_{V|S}[z^k|S]] = E_S[e^{-(1-z)\lambda S}] = G_S((1-z)\lambda)$$

- By plugging (3) in (2), we have

$$G_N(z) = \frac{(z-1)p_o G_S(\lambda(1-z))}{z - G_S(\lambda(1-z))}$$

- For M/M/1, we have  $G_S(s) = \frac{\mu}{\mu+s} \Rightarrow G_S((1-z)\lambda) = \frac{\mu}{\mu+(1-z)\lambda} = 1/(1+(1-z)\rho)$

$$\begin{aligned} G_N(z) &= \frac{(z-1)(1-\rho)}{z(1+(1-z)\rho) - 1} = \frac{1-\rho}{1-\rho z} \\ &= (1-\rho)(1+\rho z + (\rho z)^2 + \dots) \end{aligned}$$

### 3 Homework

- Consider a system with two components whose failure rate is  $\lambda$ . The component repair time is exponentially distributed with rate  $\mu$ . However, the system fails if both components fail. Find the mean time to failure of such system. Characterize the failure time of this system.